

# Biost 518 Applied Biostatistics II

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## Lecture 9: Multiple Regression: Modeling Dose - Response

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## Lecture Outline

- .....
- Modeling complex “dose response”
  - Flexible methods

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## Modeling Complex “Dose-Response”

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## Linear Predictors

- .....
- The most commonly used regression models use “linear predictors”
    - “Linear” refers to linear in the parameters
    - The modeled predictors can be transformations of the scientific measurements

- Examples

$$g[\theta | X_i, W_i] = \beta_0 + \beta_{\log X} \times \log(X_i)$$

$$g[\theta | X_i, W_i] = \beta_0 + \beta_X \times X_i + \beta_{X^2} \times X_i^2$$

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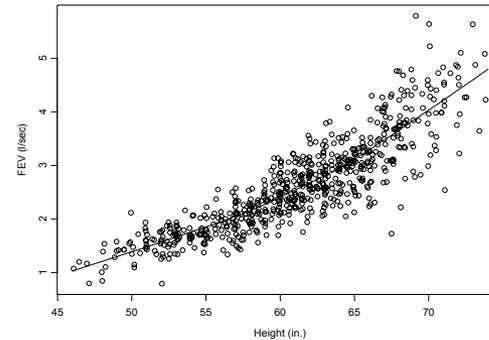
## Transformations of Predictors

- We transform predictors to provide more flexible description of complex associations between the response and some scientific measure
  - Threshold effects
  - Exponentially increasing effects
  - U-shaped functions
  - S-shaped functions
  - etc.

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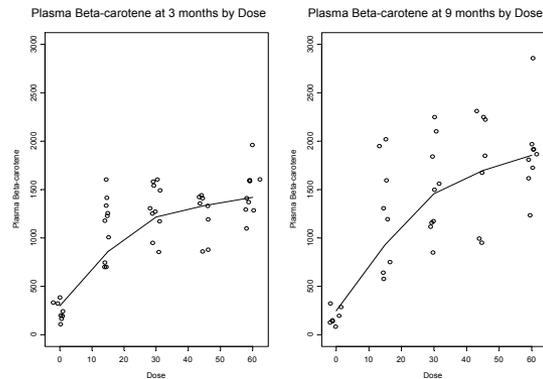
## Ex: Cubic Relationship

FEV vs Height in Children



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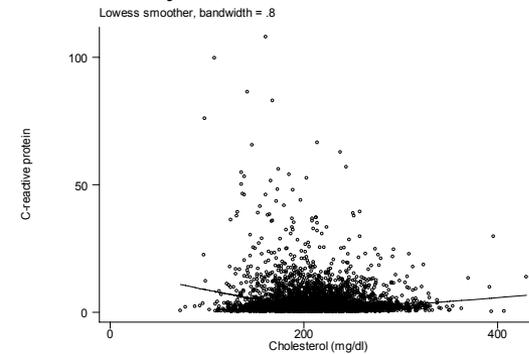
## Ex: Threshold Effect of Dose?



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## Ex: U-shaped Trend?

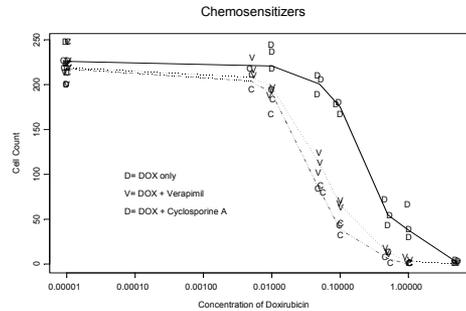
- Inflammatory marker vs cholesterol



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## Ex: S-shaped trend

- *In vitro* cytotoxic effect of Doxorubicin with chemosensitizers



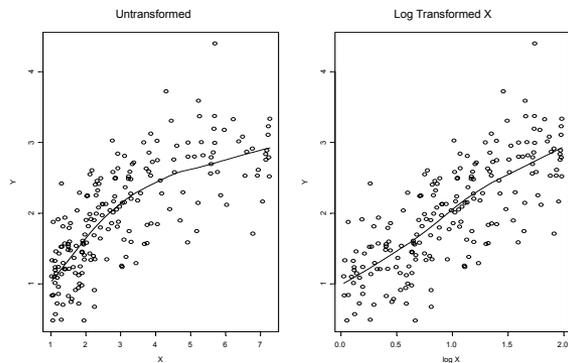
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## “1:1 Transformations”

- Sometimes we transform 1 scientific measurement into 1 modeled predictor
  - Ex: log transformation will sometimes address apparent “threshold effects”
  - Ex: cubing height produces more linear association with FEV

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## Log Transformations



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## “1:Many Transformations”

- Sometimes we transform 1 scientific measurement into several modeled predictor
  - Ex: “polynomial regression”
  - Ex: “dummy variables” (“factored variables”)
  - Ex: “piecewise linear”
  - Ex: “splines”

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## Polynomial Regression

- Fit linear term plus higher order terms (squared, cubic, ...)
- Can fit arbitrarily complex functions
  - An n-th order polynomial can fit n+1 points exactly
- Generally very difficult to interpret parameters
  - I usually graph function when I want an interpretation
- Special uses
  - 2<sup>nd</sup> order (quadratic) model to look for U-shaped trend
  - Test for linearity by testing that all higher order terms have parameters equal to zero

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## Ex: FEV – Height Assoc Linear?

- We can try to assess whether any association between mean FEV and height follows a straight line association
  - I fit a 3<sup>rd</sup> order (cubic) polynomial due to the known scientific relationship between volume and height

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## Ex: FEV – Height Assoc Linear?

```

. g htsqr= height^2
. g htcub = height^3
. regress fev height htsqr htcub, robust
Linear regression          Number of obs =      654
                          Prob > F          =  0.0000
                          R-squared          =  0.7742
                          Root MSE       =  .41299
  
```

	Robust					
fev	Coef	SE	t	P> t	[95% C I]	
height	.0306	.635	0.05	0.962	-1.22	1.28
htsqr	-.0015	.0108	-0.14	0.888	-.0227	.0196
htcub	.00003	.00006	0.43	0.671	-.00009	.0001
_cons	.457	12.4	0.04	0.971	-23.8	24.76

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## Ex: FEV – Height Assoc Linear?

- Note that the P values for each term were not significant
  - But these are addressing irrelevant questions:
    - After adjusting for 2<sup>nd</sup> and 3<sup>rd</sup> order relationships, is the linear term important?
    - After adjusting for linear and 3<sup>rd</sup> order relationships, is the squared term important?
    - After adjusting for linear and 2<sup>nd</sup> order relationships, is the cubed term important?
  - We need to test 2<sup>nd</sup> and 3<sup>rd</sup> order terms simultaneously

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## Ex: FEV – Height Assoc Linear?

```
.....  
. test htsqr htcub  
  
( 1) htsqr = 0  
( 2) htcub = 0  
  
F( 2, 650) = 30.45  
Prob > F = 0.0000
```

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## Ex: FEV – Height Assoc Linear?

- ```
.....
```
- We find clear evidence that the trend in mean FEV versus height is nonlinear
    - (Had we seen  $P > 0.05$ , we could not be sure it was linear– it could have been nonlinear in a way that a cubic polynomial could not detect)

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## Ex: log FEV – Ht Assoc Linear?

- ```
.....
```
- We can try to assess whether any association between mean log FEV and height follows a straight line association
    - I again fit a 3<sup>rd</sup> order (cubic) polynomial, but don't really have a good reason to do this rather than some other polynomial

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## Ex: log FEV – Ht Assoc Linear?

```
.....  
. g logfev = log(fev)  
. regress logfev height htsqr htcub, robust  
Linear regression      Number of obs =      654  
                      F( 3, 650) = 730.53  
                      Prob > F    = 0.0000  
                      R-squared    = 0.7958  
                      Root MSE   = .15094  
  
_____+-----  
|                Robust  
logfev |   Coef   SE      t  P>|t|  [95% C I]  
height |   .0707  .24835  0.28  0.776  -.417   .558  
htsqr  |  -.0002  .00410 -0.04  0.964  -.0082  .008  
htcub  |  3e-07   .00002  0.01  0.989  -.00004 .00004  
_cons  |  -2.79   4.985  -0.56  0.576  -12.6   6.997
```

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## Ex: log FEV – Ht Assoc Linear?

.....

- Note that again that the P values for each term were not significant
  - But these are addressing irrelevant questions:
  - We need to test 2<sup>nd</sup> and 3<sup>rd</sup> order terms simultaneously

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## Ex: log FEV – Ht Assoc Linear?

.....

```
. test htsqr htcub
```

```
( 1) htsqr = 0
```

```
( 2) htcub = 0
```

```
F( 2, 650) = 0.29  
Prob > F = 0.7464
```

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## Ex: log FEV – Ht Assoc Linear?

.....

- We do not find clear evidence that the trend in mean FEV versus height is nonlinear
  - This does not prove linearity, because it could have been nonlinear in a way that a cubic polynomial could not detect
    - (But I would think that the cubic would have picked up most patterns of nonlinearity likely to occur in this setting)

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## Ex: log FEV – Ht Assoc Linear?

.....

- We have not addressed the question of whether log FEV is associated with height
  - This question could have been addressed in the cubic model by
    - Testing all three height-derived variables simultaneously
    - OR (because only height-derived variables are included in the model) looking at the overall F test
  - Alternatively, fit a model with only the height
    - But generally bad to go fishing for models

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## Ex: log FEV – Ht Assoc?

```

.....
. regress logfev height, robust
Linear regression      Number of obs =    654
                      F( 1, 652) = 2155.08
                      Prob > F   = 0.0000
                      R-squared   = 0.7956
                      Root MSE  = .15078

```

	Robust					
<u>logfev</u>	<u>Coef</u>	<u>StdErr</u>	<u>t</u>	<u>P&gt; t </u>	<u>[95% CI]</u>	
height	.0521	.0011	46.42	0.000	.0499	.0543
_cons	-2.27	.0686	-33.13	0.000	-2.406	-2.137

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## Dummy Variables

- Indicator variables for all but one group
  - This is the only appropriate way to model nominal (unordered) variables
    - E.g., for marital status
      - Indicator variables for
        - » married (married = 1, everything else = 0)
        - » widowed (widowed = 1, everything else = 0)
        - » divorced (divorced = 1, everything else = 0)
        - » (single would then be the intercept)
  - Often used for other settings as well
  - Equivalent to “Analysis of Variance (ANOVA)”<sup>26</sup>

## Ex: Mean Salary by Field

- Field is a nominal variable, so we must use dummy variables
  - I decide to use “Other” as a reference group, so generate new indicator variables for Fine Arts and Professional fields

```

.....
. g arts= 0
. replace arts=1 if field==1
(2840 real changes made)
. g prof= 0
. replace prof=1 if field==3
(3809 real changes made)

```

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## Ex: Mean Salary by Field

```

.....
. regress salary arts prof if year==95, robust
Linear regression      Number of obs =   1597
                      F( 2, 1594) = 120.85
                      Prob > F   = 0.0000
                      R-squared   = 0.1021
                      Root MSE  = 1931.2

```

	Robust					
<u>salary</u>	<u>Coef</u>	<u>SE</u>	<u>t</u>	<u>P&gt; t </u>	<u>[95% CI]</u>	
arts	-1014	105	-9.67	0.000	-1219	-808
prof	1225	134	9.16	0.000	963	1487
_cons	6292	61.1	103.03	0.000	6172	6411

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## Ex: Interpretation of Intercept

.....

- Based on coding used
  - Intercept corresponds to mean salary for faculty in “Other” fields
    - These faculty will have arts==0 and prof==0
  - Estimated mean salary is \$6,292 / month
  - 95% CI: \$6,172 to \$6,411 / month
  - Highly statistically different from \$0 / month

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## Ex: Interpretation of Slopes

.....

- Based on coding used
  - Slope for “arts” is difference in mean salary between “Fine Arts” and “Other” fields
    - Fine arts faculty will have arts==1 and prof==0; “Other” fields will have arts==0 and prof==0
  - Estimated difference in mean monthly salary is \$1,014 lower for fine arts
  - 95% CI: \$808 to \$1,219 / month lower
  - Highly statistically different from \$0

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## Ex: Interpretation of Slopes

.....

- Based on coding used
  - Slope for “prof” is difference in mean salary between “Professional” and “Other” fields
    - Professional faculty will have arts==0 and prof==1; “Other” fields will have arts==0 and prof==0
  - Estimated difference in mean monthly salary is \$1,225 higher for professional
  - 95% CI: \$963 to \$1,487 / month higher
  - Highly statistically different from \$0

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## Ex: Descriptive Statistics

.....

- Because we modeled the three groups with two predictors plus intercept, the estimates agree exactly with sample means

```
. table field if year==95, co(mean salary)
```

<u>field</u>	<u>mean(salary)</u>
Arts	5278.082
Other	6291.638
Prof	7516.67

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## Stata: "Predicted Values"

- After computing a regression model, Stata will provide "predicted values" for each case
  - Covariates times regression parameter estimates for each case
  - `"predict varname"`

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## Ex: Salary by Field

```
. predict fit
(option xb assumed; fitted values)
. bysort field: summ fit
-> field = Arts
Vrbl | Obs      Mean    SD      Min      Max
fit  | 220  5278.082    0  5278.082  5278.082
-> field = Other
Vrbl | Obs      Mean    SD      Min      Max
fit  | 1067  6291.638    0  6291.638  6291.638
-> field = Prof
Vrbl | Obs      Mean    SD      Min      Max
fit  | 310   7516.67    0  7516.67   7516.67
```

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## Ex: Hypothesis Test

- To test for different mean salaries by field
    - We have modeled field with two variables
      - Both slopes would have to be zero for there to be no association between field and mean salary
    - Simultaneous test of the two slopes
      - We can use the Stata "test" command
- ```
. test arts prof
      F( 2, 1594) = 120.85
      Prob > F = 0.0000
```
- OR because only field variables are in the model, we can use the overall F test

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## Stata: Dummy Variables

- Stata has a facility to automatically create dummy variables
  - Prefix regression commands with `"xi: ..."`
  - Prefix variables to be modeled as dummy variables with `"i.varname"`
  - (Stata will drop the lowest category)

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## Stata: Dummy Variables

```
.....
. xi: regress salary i.field if year==95, robust
i.field _ifield_1-3(ntrllly coded; _ifield_1 omitted)
Linear regression      Number of obs =   1597
                      F( 2, 1594) = 120.85
                      Prob > F   = 0.0000
                      R-squared   = 0.1021
                      Root MSE  = 1931.2
```

|           | Robust |      |       |       |           |      |
|-----------|--------|------|-------|-------|-----------|------|
| salary    | Coef   | SE   | t     | P> t  | [95% C I] |      |
| _ifield_2 | 1014   | 105  | 9.67  | 0.000 | 808       | 1219 |
| _ifield_3 | 2239   | 146  | 15.30 | 0.000 | 1952      | 2526 |
| _cons     | 5278   | 85.2 | 61.94 | 0.000 | 5111      | 5445 |

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## Ex: Correspondence

- ```
.....
```
- This regression model is the exact same as the one in which I modeled “arts” and “prof”
    - Merely “parameterized” (coded) differently
  - Two models are equivalent if they lead to the exact same estimated parameters
    - Inference about corresponding parameters will be the same no matter how it is parameterized

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## Continuous Variables

- ```
.....
```
- We can also use dummy variables to represent continuous variables
    - Continuous variables measured at discrete levels
      - E.g., dose in an interventional experiment
    - Continuous variables divided into categories

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## Relative Advantages

- ```
.....
```
- Dummy variables fits groups exactly
    - If no other predictors in the model, parameter estimates correspond exactly with descriptive statistics
  - With continuous variables, dummy variables assume a “step function” is true
  - Modeling with dummy variables ignores order of predictor of interest

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## Flexible Methods

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## Flexible Modeling of Predictors

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- We do have methods that can fit a wide variety of curve shapes
  - Dummy variables
    - A step function with tiny steps
  - Polynomials
    - If high degree: allows many patterns of curvature
  - Splines
    - Piecewise linear or piecewise polynomial
  - Fractional polynomial

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## Stata: Linear Splines

.....

- Stata will make variable that will fit piecewise linear curves
  - Joined at “knots”
  - Lines in between
- `mkspline newvar0 #k1 newvar1 #k2 newvar2 ... #kp varp= oldvar`
  - Regression on `newvar0 ... newvarp`
    - Straight lines between min and k1; k1 and k2, etc.

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## Ex: Height vs Age in Children

.....

```
. mkspline age6A 9.5 age12A 15.5 age17A=age

. regress height age6A age12A age17A
height |      Coef.   Std. Err.      t    P>|t|     [95% Conf. Interval]
age6A |  2.488554   .0999399    24.90  0.000   2.292311   2.684798
age12A |  1.085214   .088609    12.25  0.000   .9112196   1.259209
age17A | -.5530841   .3713509    -1.49  0.137  -1.282276   .176108
   _cons | 38.49107   .8131791    47.33  0.000  36.8943   40.08785

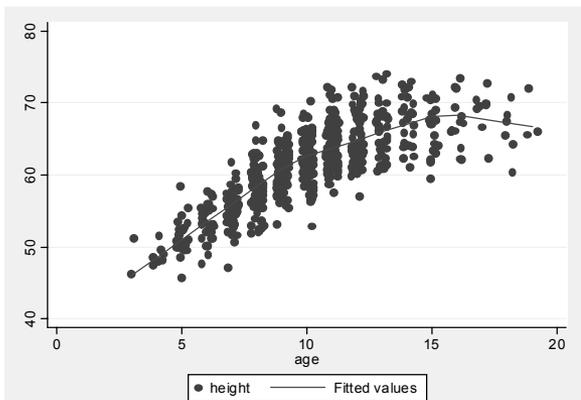
. predict fitA

. sort age

. twoway (scatter height age, jitter(3)) (line fitA
age)
```

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## Ex: Height vs Age: 2 Knots



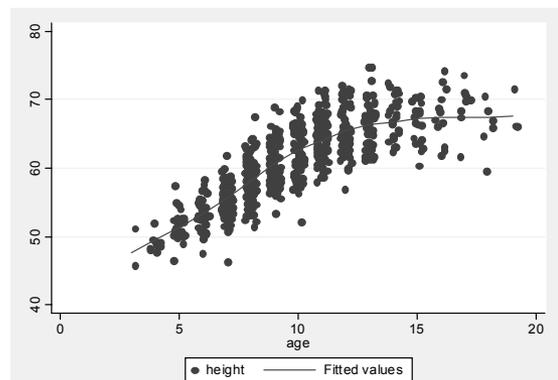
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## Ex: Height vs Age: 4 Knots

```
.....  
. mkspline age4 6.5 age8 9.5 age11 12.5 age14 15.5  
  age17=age  
  
. regress height age4 age8 age11 age14 age17  
  
. predict fit  
  
. sort age  
  
. twoway (scatter height age, jitter(3)) (line fit  
  age)
```

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## Ex: Height vs Age: 4 Knots



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## Stata: fracpoly

- ```
.....
```
- Stata will make variables modeling “fractional polynomials”
    - Can fit many different shapes depending on degree of the fractional polynomials
    - Can ask Stata to find “best” degree of the fractional polynomials: “fracpoly”
    - Can ask Stata to make new variables to model fractional polynomial of desired degree: “fracgen”

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## fracpoly

- Command

```
fracpoly regressioncommand yvar xvar,  
degree (#)
```

### Example

```
fracpoly regression logslry yrdeg,  
degree (3)
```

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## Ex: 3<sup>rd</sup> Degree fracpoly

```
-> gen double Iyrde__1 = X^3-374.9107994 if  
e(sample)  
-> gen double Iyrde__2 = X^3*ln(X)-740.6597949 if  
e(sample)  
-> gen double Iyrde__3 = X^3*ln(X)^2-1463.219871 if  
e(sample)  
(where: X = yrdeg/10)
```

(Regression output omitted)

Deviance: 21329.66.

Best powers of yrdeg among 164 models fit: 3 3 3.

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## Adjusting for Confounding

```
. regress logslry female Iyrde__1 Iyrde__2 Iyrde__3,  
robust
```

|                | Robust      |           |          |                 |                  |       |
|----------------|-------------|-----------|----------|-----------------|------------------|-------|
| <u>logslry</u> | <u>Coef</u> | <u>SE</u> | <u>t</u> | <u>P&gt; t </u> | <u>[95% C I]</u> |       |
| female         | -.119       | .007      | -16.87   | 0.000           | -.133            | -.106 |
| Iyrde__1       | -.111       | .014      | -7.68    | 0.000           | -.139            | -.082 |
| Iyrde__2       | .088        | .012      | 7.30     | 0.000           | .065             | .112  |
| Iyrde__3       | -.018       | .003      | -6.91    | 0.000           | -.023            | -.013 |
| _cons          | 8.36        | .004      | 1983.98  | 0.000           | 8.35             | 8.36  |

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