

**Biost 517**  
**Applied Biostatistics I**  
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Lecture 14:  
Two Sample Inference with Right Censored  
Data

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**Lecture Outline**  
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- Comparing Independent Proportions
  - Large Samples (Censored)
    - Using Kaplan-Meier Estimates
- Comparing Hazard Functions
  - Logrank Test
  - Wilcoxon Test for Censored Data
- Comparing Quantiles
  - Parametric Accelerated Failure Time Models

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**Right Censored Data**  
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**Right Censored Data**  
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- Recall from Lecture 6: Censored variables
  - A special type of missing data (the exact value is not always known)
    - Right censoring: for some observations it is only known that the true value exceeds some threshold
    - Left censoring: for some observations it is only known that the true value is below some threshold
    - Interval censoring: for some observations it is only known that the true value is between some thresholds

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### Examples

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- PSA data set
  - Subjects were followed with serial PSAs
  - Interested in time to relapse
  - Some still in remission at time of analysis
  - (Ignoring these subjects is ignoring successes)
- University salary data set
  - Interest is in sex discrimination
  - Interested in time to promotion from associate
  - Some subjects have not yet been promoted
  - (Ignoring these subjects may be ignoring discrimination)

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### Descriptive Statistics

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- Sample mean, sample median (and other quantiles), sample standard deviation and variance are not appropriate
- Instead, descriptive statistics must be computed from Kaplan-Meier estimates
  - Only exception: You could use binomial proportions to estimate survival to the first censoring time
    - E.g., PSA data: All subjects followed at least 24 months

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### Noninformative Censoring

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- Recall: Our estimation methods only appropriate if censoring is not informative about subjects who were either more or less likely to have an event in the immediate future
  - Censored subjects must look like a random sample of those at risk at time of censoring
  - (Later we shall say that they are a random sample from all subjects at risk having similar modeled covariates)

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### Comparing Independent Proportions

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Large Samples with Right Censored Data

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## Kaplan-Meier Estimates

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- Estimate  $S(t) = Pr(T^0 > c)$  for arbitrary  $c$ 
  - Nonparametric
    - Works for all distributions
    - (Also works for uncensored data)
  - Consistent for true value in infinite samples
  - Can derive estimates of quantiles
  - Can only estimate mean if estimated survival curve goes to 0
    - But can define “restricted mean” up to some time

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## Approximate Distribution

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- If interested in  $\theta = S(c) = Pr(T^0 \geq c)$  in presence of right censoring

Kaplan - Meier estimates for  $i$ th group

$$\theta_i = \hat{S}_i(c) = \prod_{j: t_j \leq c} \left( 1 - \frac{d_{ij}}{n_{ij}} \right) \sim N \left( S_i(c), [se(\hat{S}_i(c))]^2 \right)$$

(with  $se(\hat{S}_i(c))$  from Greenwood's formula)

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## Stata: Kaplan-Meier Commands

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- Syntax for “setting survival data”
  - `“stset endtime eventind, t0(entrytime)”`
  - *endtime*: name of the variable measuring the time at the end of the interval
  - *eventind*: name of an indicator (0 or 1) variable indicating event status at the end of the interval
  - *entrytime*: name of the variable specifying the time at the start of the interval
    - (does not need to be supplied)
  - `“stset, clear”` resets the data set

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## Stata: Kaplan-Meier Commands

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- Syntax for getting estimates, plots
  - Plotting survival curves
    - `“sts graph”`
    - `“sts graph, atrisk”`
    - `“sts graph, cens(s)”`
  - Listing survival estimates
    - `“sts list”`
  - Saving survival estimates
    - `“sts gen newvar = s”`

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## Two Group Comparisons

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- To compare survival probabilities, we would compute SE for each group individually, then use methods for combining estimates

For independent  $\hat{\theta}_1 \sim N(\theta_1, se_1^2)$ ;  $\hat{\theta}_2 \sim N(\theta_2, se_2^2)$

$$\hat{\theta}_1 + \hat{\theta}_2 \sim N(\theta_1 + \theta_2, se_1^2 + se_2^2)$$

$$\hat{\theta}_1 - \hat{\theta}_2 \sim N(\theta_1 - \theta_2, se_1^2 + se_2^2)$$

$$\hat{\theta}_1 / \hat{\theta}_2 \sim N\left(\frac{\theta_1}{\theta_2}, \frac{1}{\theta_2^2} \left( se_1^2 + \frac{\theta_1^2}{\theta_2^2} se_2^2 \right)\right)$$

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## Two Group Inference

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- As with any (approximately) normally distributed estimator, CI and P values are computed using

100(1- $\alpha$ )% confidence interval

$$(est) \pm z_{1-\alpha/2} \times (std\ err)$$

Normalized Z statistic for  $H_0: \theta = \theta_0$

$$Z = \frac{(est) - (null\ hyp)}{(std\ err)}$$

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## Example: PSA Data

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- Men with prostate cancer
  - Hormonal treatment
  - Followed for signs of progression
- Interested in estimating probability of remaining in remission for three years
  - Testing hypothesis that three year survival differs between bone scan score less than 3 or bone scan score equals 3

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## Example: Preparing Data

.....

- Reading in data (note string variable)
 

```
. infile ... obstime str8 inrem using psa.txt
```
- Creating indicator of relapse
 

```
. g relapse = 0
      . replace relapse = 1 if inrem=="no"
```
- "Setting" survival variables
 

```
. stset obstime relapse
```

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### Dichotomizing Bone Scan Score

- Method 1 (must consider missing data)

```
. g bss3= 0
. replace bss3=1 if bss==3
. replace bss3=. if bss==.
```

- Method 2 (recode handles missing data)

```
. g bss3= bss
. recode bss3 1/2=0 3=1
```

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### Stata: KM Listing

```
. sts list, by(bss3) at(12 24 36 48)
Beg.          Surv   Std.
Time Total Fail   Fctn Error [95% Conf Int]
bss3=0
 12   18   1   0.9444 0.0540   0.6664 0.9920
 24   14   3   0.7778 0.0980   0.5110 0.9102
 36   12   1   0.7130 0.1092   0.4398 0.8699
 48    6   3   0.4801 0.1356   0.2101 0.7082
bss3=1
 12   22  10   0.6667 0.0861   0.4692 0.8047
 24   15   6   0.4667 0.0911   0.2839 0.6304
 36    9   5   0.2963 0.0841   0.1464 0.4630
 48    2   4   0.1058 0.0659   0.0209 0.2713
```

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### Stata: Difference and SE

- Three year survival probabilities
  - Bone scan score < 3: 0.7130 (SE 0.1092)
  - Bone scan score = 3: 0.2963 (SE 0.0841)

- Estimated diff in 3 year survival probability

```
. display 0.7130 - 0.2963
.4167
```

- Standard error of estimated difference

```
. display sqrt( 0.1092^2 + 0.0841^2 )
.13783124
```

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### Stata: 95% CI and P value

- 95% confidence interval: 0.147 to 0.687

```
. display invnorm(.975)
1.959964
. display .4167 - invnorm(.975) * 0.13783124
.14655573
. display .4167 + invnorm(.975) * 0.13783124
.68684427
```

- Two-sided P value : P = 0.0025

– (note use of negative)

```
. display 2 * normprob( - .4167/.13783124)
.00250065
```

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### Interpretation

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- The Kaplan-Meier estimate of difference in survival is that men with a bone scan score less than 3 have an absolute improved 3 year survival of 41.7% relative to bss=3
- With 95% confidence, such an observation is not consistent with a true absolute improvement less than 14.7% or greater than 68.7%
- Based on the P value of 0.0025, we reject the null hypothesis of no association between bone scan score and 3 year survival prob

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### Comparing Hazard Functions

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Logrank Test

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### Scientific Questions

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- With time to event data, we are generally interested in probability that an event will occur in a specified time
- Right censored data presented problems, because the measurement of events was over varying amounts of time
  - Effect modification by time?
  - Confounding by time?
  - Increased precision by accounting for time?

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### General Strategy

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- We want to use methods that adjust for the time of observation
- Kaplan-Meier estimates at a fixed time
- Logrank and modified Wilcoxon statistics by averaging effects over time

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## Hazard Function

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- With censored data, we often compare probability distns using hazard functions
  
- Hazard = Instantaneous risk of an event
  - Among subjects at risk of an event, what is the probability of having an event in the next instant
  
- Advantage of using hazard with censored data
  - Only need to consider subjects currently at risk
  - Only need to consider whether they have an event right then

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## Hazard Function

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- Estimates of the hazard at each time look somewhat like a binomial proportion
  
- We do not often estimate the hazard function over time
  
- However, we do compare hazard functions
  - Usually we estimate a hazard ratio: relative risk of an event
  - We want to average the estimates of the hazard ratio over all times

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## Stratified Analyses

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- Recall that we are often interested in comparing groups within strata
  - Confounding:
    - Comparisons within strata are all similar, but failure to stratify results in a comparison that is misleading due to bias
      - There are nuances here as we go from analyses of means to analyses of nonlinear summary measures (e.g., odds- more later)
  - Interactions:
    - Comparisons within strata result in different estimates

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## Adjusting for Covariates

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- We can remove confounding by “adjusting” for the confounder using a stratified test statistic
  
- “Adjustment” for a covariate means making comparisons between subjects who have similar levels of that covariate
  - E.g., in FEV data, compare smoking children to nonsmokers of same age, height
  - Average the differences seen in age, height strata

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### Role of Effect Modification

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- Adjustment for a covariate does not remove interactions
  - Interactions means that the question has different answers in different strata
- Adjustment for a covariate will merely average the effect across strata
  - Usually weighted by the sample size in each stratum

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### Stratified Estimates

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- Obtained by combining estimates from each (independent) stratum
- Generally, best to average the estimates (sometimes weighted) rather than Z scores
- SEs for the stratified estimates are obtained using properties of independent random variables
  - Standard errors are the sum of squared standard errors from the independent strata

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### Example

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- Effect of hepatomegaly on survival after adjustment for sex?
  - Summarize response by 5 year survival
  - Hepatomegaly effect by sex: For each sex, compute difference in survival across hepatomegaly groups
  - Adjusted measure of effect: Compute the average difference between hepatomegaly effects
    - Usually a weighted average

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### SE for Stratified Estimates

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	Males	Females
Hepatomegaly	$\hat{\theta}_{M1} \sim N(\theta_{M1}, se(\hat{\theta}_{M1}))$	$\hat{\theta}_{F1} \sim N(\theta_{F1}, se(\hat{\theta}_{F1}))$
No Hepatomegaly	$\hat{\theta}_{M0} \sim N(\theta_{M0}, se(\hat{\theta}_{M0}))$	$\hat{\theta}_{F0} \sim N(\theta_{F0}, se(\hat{\theta}_{F0}))$
	↓	
Weighted average	$p(\hat{\theta}_{M1} - \hat{\theta}_{M0}) - (1-p)(\hat{\theta}_{F1} - \hat{\theta}_{F0})$	
Approx Distn	$\sim N(\text{mean} = p(\theta_{M1} - \theta_{M0}) - (1-p)(\theta_{F1} - \theta_{F0}),$	
	$se = \sqrt{p^2(se^2(\hat{\theta}_{M1}) + se^2(\hat{\theta}_{M0})) + (1-p)^2(se^2(\hat{\theta}_{F1}) + se^2(\hat{\theta}_{F0}))}$	

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### Mantel-Haenszel Statistic

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- Generally regarded as the best choice of methods for comparing binary data across strata
  
- Based on the odds rather than the proportion
  - It is rare that we might expect the difference in proportions to be constant across strata
  
- Other methods can be based on the asymptotic distribution of the log odds
  - (More on these methods next quarter)

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### Logrank Test

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- The Mantel-Haenszel test is also the basis for a very popular method of comparing censored survival data across populations: The logrank statistic
  - The data are stratified by time of event
    - Often only a single event is observed in each stratum
    - Stratified estimates of the odds ratio are obtained

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### Noninformative Censoring

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- Most often the same subjects are used in several different strata
  
- Noninformative censoring argues that the estimates are independent across strata asymptotically

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### Tests Equality of Hazards

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- Equal hazard functions implies equal distributions
  - The P value for this test is interpretable as a test that the survival distributions are similar for the two groups
  
- This test is more powerful when the true alternative is “proportional hazards”
  - Proportional hazards = constant risk ratio over time
  - Proportional hazards regression will provide estimates of the risk ratio

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### Logrank Test: Stata Commands

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- The logrank test can be obtained from Stata using the "sts test" command (after defining survival variables using "stset")
  - "sts test *groupvar*, logrank"
    - *groupvar* indicates the groups to be compared
    - logrank test is default
    - P value based on a chi square statistic
      - Hence a two-sided P value
      - (Obtaining a one-sided P value is deferred until we discuss proportional hazards regression next quarter)

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### Example: PSA Survival by bss

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```
. sts test bss3
Log-rank test for equality of survivor functions
```

bss3	Events	Events
	observed	expected
0	9	17.18
1	25	16.82
<b>Total</b>	<b>34</b>	<b>34.00</b>

```
chi2(1) = 8.30
Pr>chi2 = 0.0040
```

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### Example: Interpretation

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- Based on the two-sided P value of 0.004, we reject the null hypothesis of equal relapse free survival probabilities between the bone scan score groups
  - (Because the expected events are less than observed in the bss=3 group, we can presume that the higher bss is associated with worse relapse free survival)

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### Hazard Ratio Estimates

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- Logrank test does not give estimates
  - However, it is closely related to "proportional hazards regression" ("Cox regression")
    - Provides estimates of the (average) hazard ratio over time
- Hazard ratio
  - Groups with higher hazards have higher event rates
    - Hazard ratio greater than 1 = Worse "survival"

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## Proportional Hazard Regression

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- HR estimates approximately normal in large samples
- Stata commands
  - "stcox groupvar, robust"
    - "robust" eliminates need for proportional hazards
    - Gives hazard ratio, 95% CI
      - CI is computed on log hazard ratio scale
    - P values
      - "Wald test" (based on approximately normal estimate)
      - "Likelihood ratio test"
      - ("Score test" would be the logrank test)

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## Example: PSA Survival by bss

.....

```
. stcox bss3, robust
No. of subjects = 48      Number of obs =    48
No. of failures = 34      Time at risk =   1408
                               LR chi2(1) =    8.35
Log likelihood = -106.9   Prob > chi2 = 0.0038
```

Robust						
t	HazRat	StdErr	z	P> z	[95% CI]	
bss3	2.96	1.11	2.89	0.004	1.42	6.16

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## Example: Interpretation

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- We estimate that at any given time the risk of relapse in men with bss=3 tends to be 2.96 times that of men with lower bss
- 95% CI suggests these results typical if true risk of relapse with bss=3 is 1.42 to 6.16 times that in men with lower bss
- Based on P value of 0.004 we would reject null hypothesis of no association between relapse and bss

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## Comparing Quantiles

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Parametric Models  
for Censored Data

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### Parametric Models for Censored Data

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- There are times that inference for censored data is based on parametric models
  - Accelerated failure time models
    - Assume a constant ratio between groups for all quantiles of survivor distribution
    - E.g., dogs live 7 years for each year of human life

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### Parametric Models for Censored Data

.....

- Commonly used parametric models
  - Exponential:
    - Constant hazard independent of past
  - Weibull:
    - Theoretical derivation: First failure in a series of components (weakest link in a chain)
    - Log hazard is linear
    - Exponential is special case
    - Only accelerated failure time model that is also proportional hazards

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### Parametric Models for Censored Data

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- Commonly used parametric models (cont.)
  - Gamma:
    - Theoretical derivation: Final failure in parallel components
    - Exponential is special case
  - Lognormal
  - Many other generalizations

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### Caveats

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Choice of Summary Measures

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### Parametric Models

- All of the parametric models will be sensitive to violation of the distributional assumptions
  - Because these models assume constant ratio of all quantiles, we do not have robustness to other distributions in any particular model (including lognormal)
  - (We will discuss these models with regression next quarter)

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### Semiparametric Models

- We do know how to use the proportional hazards model, even when the hazard ratio is not constant
  - However, you need to be careful– it may not estimate anything you care about

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### Hypothetical Example: Setting

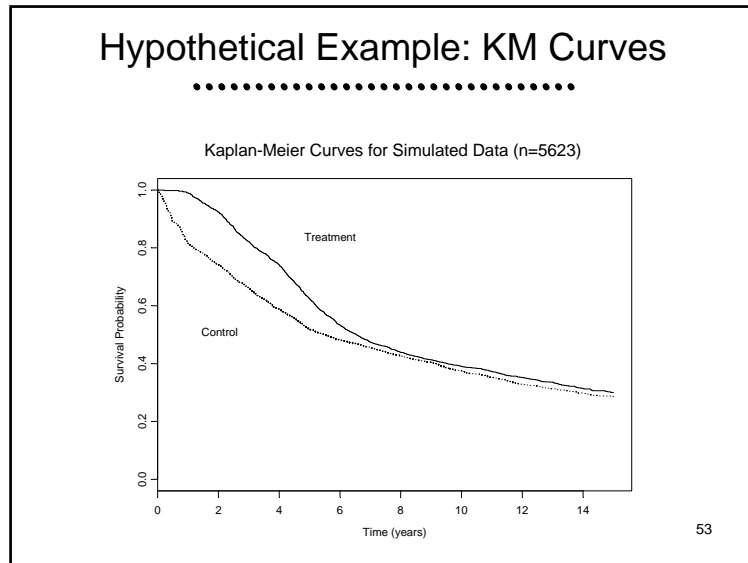
- Consider survival with a particular treatment used in renal dialysis patients
- Extract data from registry of dialysis patients
- To ensure quality, only use data after 1995
  - Incident cases in 1995: Follow-up 1995 – 2002 (8 yrs)
  - Prevalent cases in 1995: Data from 1995 - 2002
    - Incident in 1994: Information about 2<sup>nd</sup> – 9<sup>th</sup> year
    - Incident in 1993: Information about 3<sup>rd</sup> – 10<sup>th</sup> year
    - ...
    - Incident in 1988: Information about 8<sup>th</sup> – 15<sup>th</sup> year

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### Hypothetical Example: Analysis

- Choice of summary measure
  - Survival at fixed point in time
  - Median, other quantiles
  - Mean (or restricted mean)
  - Hazard ratio (or weighted average of hazard ratio over time)
- Choice of methods
  - Parametric, semiparametric, nonparametric

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### Who Wants To Be A Millionaire?

.....

Proportional hazards analysis estimates a **Treatment** :  
**Control** hazard ratio of

- A: 2.07 (logrank P = .0018)
- B: 1.13 (logrank P = .0018)
- C: 0.87 (logrank P = .0018)
- D: 0.48 (logrank P = .0018)

- Lifelines:
  - 50-50? Ask the audience? Call a friend?

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### Who Wants To Be A Millionaire?

.....

Proportional hazards analysis estimates a **Treatment** :  
**Control** hazard ratio of

- B: 1.13 (logrank P = .0018)
- C: 0.87 (logrank P = .0018)

- Lifelines:
  - 50-50? Ask the audience? Call a friend?

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