

Biost 517
Applied Biostatistics I
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Lecture 18:
**Extension to Other Simple
Regression Models**

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Lecture Outline
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- General Simple Regression Model
- Simple Logistic Regression
- Simple Proportional Hazards Regression

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**General Regression
Model**
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Types of Variables
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- Binary data
 - E.g., sex, death
- Nominal data: unordered, categorical data
 - E.g., race, marital status
- Ordinal categorical data
 - E.g., stage of disease
- Quantitative data
 - E.g., age, blood pressure
- Right censored data
 - E.g., time to death (when not everyone has died) ⁴

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Summary Measures

- The measures commonly used to summarize and compare distributions vary according to the types of data
 - Means: binary; quantitative
 - Medians: ordered; quantitative; censored
 - Proportions: binary; nominal
 - Odds: binary; nominal
 - Hazards: censored
 - hazard = instantaneous rate of failure

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Regression Models

- According to the parameter compared across groups
 - Means → Linear regression
 - Geom Means → Linear regression on logs
 - Odds → Logistic regression
 - Rates → Poisson regression
 - Hazards → Proportional Hazards regr
 - Quantiles → Parametric survival regr

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General Regression

- General notation for variables and parameter
 - Y_i Response measured on the i th subject
 - X_i Value of the predictor for the i th subject
 - θ_i Parameter of distribution of Y_i
- The parameter might be the mean, geometric mean, odds, rate, instantaneous risk of an event (hazard), etc.

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Simple Regression

- General notation for simple regression model
 - $g(\theta_i) = \beta_0 + \beta_1 \times X_i$
 - $g(\)$ "link" function used for modeling
 - β_0 "Intercept"
 - β_1 "Slope (for predictor X)"
- The link function is usually either none (means) or log (geom mean, odds, hazard)

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Borrowing Information

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- Use other groups to make estimates in groups with sparse data
 - Intuitively: 67 and 69 year olds would provide some relevant information about 68 year olds
 - Assuming straight line relationship tells us how to adjust data from other (even more distant) age groups
 - If we do not know about the exact functional relationship, we might want to borrow information only close to each group
 - (Next quarter: splines)

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Defining “Contrasts”

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- Define a comparison across groups to use when answering scientific question
 - If straight line relationship in parameter, slope is difference in parameter between groups differing by 1 year in X
 - If nonlinear relationship in parameter, slope is average difference in parameter between groups differing by 1 year in X
 - Statistical jargon: a “contrast” across the groups

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Comparison of Methods

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- The major difference between regression models is interpretation of the parameters
 - Summary: Mean, geometric mean, odds, hazards
 - Comparison of groups: Difference, ratio
- Issues related to inclusion of covariates remain the same
 - Address the scientific question
 - Predictor of interest; Effect modifiers
 - Address confounding
 - Increase precision

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Simple Logistic Regression

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Inference About the Odds

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Logistic Regression

- Binary response variable
- Allows continuous (or multiple) grouping variables
 - But is OK with binary grouping variable also
- Compares odds of response across groups
 - “Odds ratio”

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Binary Response

- When using regression with binary response variables, we typically model the (log) odds using logistic regression
 - Conceptually, there should be no problem modeling the proportion (which is the mean of the distribution)
 - However, there are several technical reasons why we do not use linear regression very often with binary response

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Why not Linear Regression?

- Many misconceptions about the advantages and disadvantages of analyzing the odds
- Reasons that I consider valid
 - Scientific basis
 - Use of odds ratios in case-control studies
 - Plausibility of linear trends and no effect modifiers
 - Statistical basis
 - Mean variance relationship (if not using robust SE)

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Science: Case-Control Studies

- Scientific interest:
 - Distribution of “effect” across groups defined by “cause”
- Common sampling schemes
 - Cohort study: Sample by exposure
 - Estimate distribution of “effect” in exposure groups
 - Case-control study: Sample by outcomes
 - Estimate distribution of exposure in outcome groups
 - E.g., proportion (or odds) of smokers among people with or without cancer

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Science: Case-Control Studies

- Estimable odds ratios for each sampling scheme
 - Cohort study
 - Odds of cancer among smokers : odds of cancer among nonsmokers
 - Case-control study
 - Odds of smoking among cancer : odds of smoking among noncancer
- Mathematically, the two odds ratios are the same

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Science: Case-Control Studies

- The odds ratio is easily interpreted when trying to investigate rare events
 - Odds = $\text{prob} / (1 - \text{prob})$
 - Rare event: $(1 - \text{prob})$ is approximately 1
 - Odds is approximately the probability
 - Odds ratio is approximately the risk ratio
 - Risk ratios are easily understood
- Case-control studies typically used when events are rare

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Science: Linearity

- Proportions have to be between 0 and 1
 - It is thus unlikely that a straight line relationship would exist between a proportion and any predictor
 - UNLESS the predictor itself is bounded
 - OTHERWISE there eventually must be a threshold above which the probability does not increase (or only increases a little)

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Science: Effect Modification

- The restriction on ranges for probabilities also make it likely that effect modification will often be present with proportions
 - Ex: 2 Yr Relapse rates by NadirPSA>4, BSS
 - If bone scan score < 3: A difference of 0.60
 - 40% of men with nadir PSA < 4 relapse in 24 months
 - 100% of men with nadir PSA > 4 relapse in 24 months
 - If bone scan score > 3:
 - 71% of men with nadir PSA < 4 relapse in 24 months
 - Thus impossible for men with nadir PSA > 4 to have an absolute difference of 0.60 higher

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Why use the odds?

- The odds of an event are between 0 and infinity
 - Recall odds = prob / (1 – prob)
 - (Even better: log (odds) are between negative infinity and positive infinity)
 - Thus, there is a greater chance that linear relationships might hold without effect modification

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Statistics: Mean-Variance

- Classical linear regression requires equal variances in each predictor group
 - With binary data, the variance within a group depends on the mean
 - For binary Y
 - E(Y) = p
 - Var (Y) = p(1 – p)
 - (With robust regression techniques, this problem not a limitation)

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Simple Logistic Regression

- Modeling odds of binary response Y on predictor X

Distribution $\Pr(Y_i = 1) = p_i$

Model $\text{logit}(p_i) = \log\left(\frac{p_i}{1-p_i}\right) = \beta_0 + \beta_1 \times X_i$

$X_i = 0$ log odds = β_0

$X_i = x$ log odds = $\beta_0 + \beta_1 \times x$

$X_i = x+1$ log odds = $\beta_0 + \beta_1 \times x + \beta_1$

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Interpretation as Odds

- Exponentiation of regression parameters

Distribution $\Pr(Y_i = 1) = p_i$

Model $\left(\frac{p_i}{1-p_i}\right) = e^{\beta_0} \times e^{\beta_1 \times X_i}$

$X_i = 0$ odds = e^{β_0}

$X_i = x$ odds = $e^{\beta_0} \times e^{\beta_1 \times x}$

$X_i = x+1$ odds = $e^{\beta_0} \times e^{\beta_1 \times x} \times e^{\beta_1}$

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Estimating Proportions

- Proportion = odds / (1 + odds)

Distribution $\Pr(Y_i = 1) = p_i$

Model
$$p_i = \frac{e^{\beta_0} \times e^{\beta_1 \times X_i}}{1 + e^{\beta_0} \times e^{\beta_1 \times X_i}}$$

$X_i = 0$
$$p_i = e^{\beta_0} / (1 + e^{\beta_0})$$

$X_i = x$
$$p_i = \frac{e^{\beta_0} \times e^{\beta_1 \times x}}{1 + e^{\beta_0} \times e^{\beta_1 \times x}}$$

$X_i = x+1$
$$p_i = \frac{e^{\beta_0} \times e^{\beta_1 \times x} \times e^{\beta_1}}{1 + e^{\beta_0} \times e^{\beta_1 \times x} \times e^{\beta_1}}$$

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Simple Logistic Regression

- Interpretation of the model
 - Odds when predictor is 0
 - Found by exponentiation of the intercept from the logistic regression: $\exp(\beta_0)$
 - Odds ratio between groups differing in the value of the predictor by 1 unit
 - Found by exponentiation of the slope from the logistic regression: $\exp(\beta_1)$

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Stata

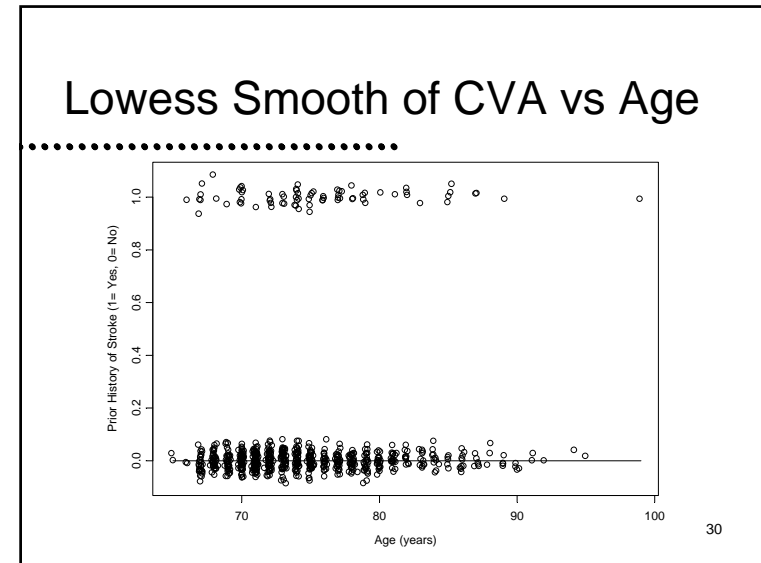
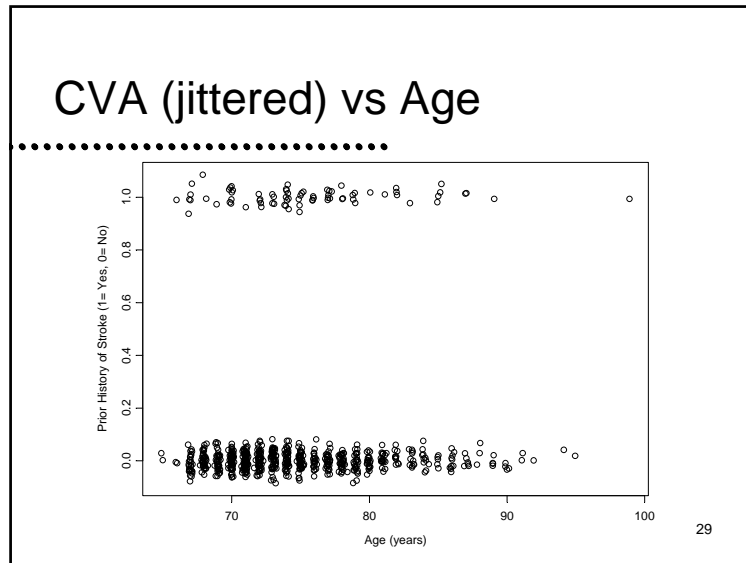
- `"logit respvar predvar, [robust]`
 - Provides regression parameter estimates and inference on the log odds scale
 - Intercept, slope with SE, CI, P values
- `"logistic respvar predvar, [robust]`
 - Provides regression parameter estimates and inference on the odds ratio scale
 - Only slope with SE, CI, P values

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Example

- Prevalence of stroke (cerebrovascular accident- CVA) by age in subset of Cardiovascular Health Study
 - Response variable is CVA
 - Binary variable: 0= no history of prior stroke, 1= prior history of stroke
 - Predictor variable is Age
 - Continuous predictor

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Characterization of Plot

- Clearly the scatterplot (even with superimposed smooth) is pretty useless with a binary response
 - (Note that we are estimating proportions– not odds– with this plot, so we can not even judge linearity for logistic regression)

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Example: Regression Model

- Answer question by assessing linear trends in log odds of stroke by age
 - Estimate best fitting line to log odds of CVA within age groups
$$\text{logodds}(CVA | Age) = \beta_0 + \beta_1 \times Age$$
 - An association will exist if the slope (β_1) is nonzero
 - In that case, the odds (and probability) of CVA will be different across different age groups

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Parameter Estimates

```

.....
. logit cva age
(iteration info deleted)

                Number of obs   =       735
                LR chi2(1)      =         2.45
                Prob > chi2     =        0.1175
                Log likelihood   = -240.98969
                Pseudo R2       =         0.0051

-----+-----
   cva |   Coef   StdErr     z    P>|z|   [95% Conf Int]
-----+-----
   age |   .0336   .0210    1.59   0.111   - .0077   .0748
   _cons |  -4.69   1.591   -2.95   0.003   -7.810  -1.572

```

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Interpretation of Stata Output

- Regression model for CVA on age
 - Intercept is labeled by “_cons”
 - Estimated intercept: -4.69
 - Slope is labeled by variable name: “age”
 - Estimated slope: 0.0336
 - Estimated linear relationship:
 - log odds relapse by nadir given by
- $$\text{log odds CVA} = -4.69 + 0.0336 \times \text{Age}_i$$

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Interpretation of Intercept

$$\text{log odds CVA} = -4.69 + 0.0336 \times \text{Age}_i$$

- Estimated log odds CVA for newborns is -4.69
 - Odds of CVA for newborns is $e^{-4.69} = 0.0092$
 - Probability of CVA for newborns
 - Use prob = odds / (1+odds): $.0092 / 1+.0092 = .0091$
- Pretty ridiculous to try to estimate
 - We never sampled anyone less than 67
 - In this problem, the intercept is just a tool in fitting the model

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Interpretation of Slope

$$\text{log odds CVA} = -4.69 + 0.0336 \times \text{Age}_i$$

- Estimated difference in log odds CVA for two groups differing by one year in age is 0.0336, with older group tending to higher log odds
 - Odds Ratio: $e^{0.0336} = 1.034$
 - For 5 year age difference: $e^{5 \times 0.0336} = 1.034^5 = 1.183$
- (If a straight line relationship is not true, we interpret the slope as an average difference in log odds CVA per one year difference in age)₃₆

Stata: “logit” versus “logistic”

- Given that we are rarely interested in the intercept, we might as well use the “logistic” command
 - It will provide inference for the odds ratio
 - We don't have to exponentiate the slope estimate

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Odds Ratios using “logistic”

```

.....
.logistic cva age
Logistic regression   Number of obs   =       735
                     LR chi2(1)           =         2.45
                     Prob > chi2         =       0.1175
                     Log likelihood      = -240.98969
                     Pseudo R2          =       0.0051

```

cva	Odds Ratio	StdErr	z	P> z	[95% Conf Int]
age	1.034	.0218	1.59	0.111	.992 1.078

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Comments on Interpretation

- I express this as a difference between group means rather than a change with aging
 - We did not do a longitudinal study
- To the extent that the true group log odds have a linear relationship, this interpretation applies exactly
 - If the true relationship is nonlinear
 - The slope estimates the “first order trend” for the sampled age distribution
 - We should not regard the estimates of individual group probabilities/odds as accurate

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Signal and Noise

- Note that the Signal and Noise idea does not apply so well here
 - We do not tend to quantify an “error distribution” with logistic regression

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Statistical Validity of Inference

- Inference (CI, P vals) about associations requires three general assumptions
 - Assumptions about approximate normal distribution for parameter estimates
 - Assumptions about independence of observations
 - Assumptions about variance of observations within groups

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Normally Distributed Estimates

- Assumptions about approximate normal distribution for parameter estimates
 - Classically or Robust SE:
 - Large sample sizes
 - Definition of “large” depends on underlying probability (odds)
 - Recall rule of thumb for chi-squared test based on expected number of events

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Independence / Dependence

- Assumptions about independence of observations for linear regression
 - Classically:
 - All observations are independent
 - Robust standard error estimates:
 - Allow correlated observations within identified clusters

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Within Group Variance

- Assumptions about variance of response within groups for logistic regression
 - Classically:
 - Mean variance relationship for binary data
 - Classical logistic regression estimates SE using model based estimates
 - Hence in order to satisfy this requirement, linearity of log odds across groups must hold
 - Robust standard error estimates:
 - Allow unequal variances across groups
 - (Do not need the linearity of log odds)

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Statistical Validity of Inference

- Inference (CI, P values) about odds of response in specific groups requires a further assumption
 - Assumption about adequacy of linear model

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Linearity of Model

- Assumption about adequacy of linear model for prediction of group odds of response with logistic regression
 - Classically OR robust standard error estimates:
 - The log odds response in groups is linear in the modeled predictor
 - (We can model transformations of the measured predictor)

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Statistical Validity of Inference

- Inference (prediction intervals, P values) about individual observations requires no further assumptions because we have binary data
 - If we know the mean (proportion), we know everything

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Implications for Inference

- Regression based inference about associations is far more robust than estimation of group odds of response
 - A hierarchy of null hypotheses
 - Strong (and intermediate) null: Total independence of Y and X
 - A binary distribution only depends on the mean (proportion, odds)
 - Weak null: No linear trend in mean of Y across X groups

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Under Strong Null

- If the response and predictor of interest were totally independent:
 - Probability of response, and hence the odds and log odds, would be the same in all groups
 - A flat line would describe the log odds response across groups (and a linear model is correct)
 - Slope would be zero
 - Within group variance would be correctly estimated by the model
 - In large sample sizes, the regression parameters are normally distributed

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Under Weak Null

- Linear trend in means across predictor groups would lie on a flat line
 - Slope of best fitting line would be zero
 - Within group variance could vary from that predicted by model
 - In large sample sizes, the regression parameters are normally distributed
 - Definition of “large” will also depend upon how much the error distributions differ across groups relative to the number sampled in each group

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Classical Logistic Regression

- Inference about slope tests strong null
 - Tests make inference assuming the null
 - The data can appear nonlinear in log odds
 - Merely evidence strong null is not true
 - Limitations
 - We cannot be confident that there is a trend in the log odds across groups
 - Valid inference about trend demands correct model
 - We cannot be confident of estimates of group probabilities (odds)
 - Valid estimates of group means demands correct model^{F1}

Robust Standard Errors

- Inference about slope tests weak null
 - Data can appear nonlinear in log odds
 - Robust SE estimates true variability
 - Does not use model based estimates of SE
 - Nonlinearity decreases precision, but inference still valid about first order (linear) trends
 - Only if linear relationship holds can we
 - Estimate group response probabilities (odds)

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Choice of Inference

- Which inference is correct?
 - Classical logistic regression and robust standard error estimates differ in the strength of necessary assumptions
 - As a rule, if all the assumptions of classical logistic regression hold, it will be more precise
 - (Hence, we will have greatest precision to detect associations if the linear model is correct)
 - The robust standard error estimates are, however, valid for detection of associations even in those instances

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Implications for Inference

- Inference about associations is far more trustworthy than estimation of group means or individual predictions
 - Nonzero slope suggests an association between response and predictor
 - Inference about linear trends in log odds if use robust SE

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Interpreting “Positive” Results

- If slope is statistically significant different from 0 using robust SE
 - Observed data is atypical of a setting with no linear trend in odds of response across groups
 - Data suggests evidence of a trend toward larger (smaller) odds in groups having larger values of the predictor
 - (To the extent the data appears linear, estimates of the group odds will be reliable)

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Interpreting “Negative” Studies

- “Differential diagnosis” of reasons for not rejecting null hypothesis of zero slope
 - There may be no association
 - [*There may be an association but not in the parameter considered (i.e., the odds of response)*]
 - There may be an association, but the best fitting line has a zero slope (a curvilinear association in the parameter)
 - There may be a first order trend in the log odds, but we lacked statistical precision to be confident that it truly exists (type II error)

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Logistic Regression Inference

- The regression output provides
 - Estimates
 - Intercept: estimated log odds CVA when age = 0
 - Slope: estimated difference in log odds CVA for two groups differing by one year in age
 - Standard errors
 - Confidence intervals
 - P values testing for
 - Intercept= zero (odds= 1; prob= 0.5) (who cares?)
 - Slope= zero (test for linear trend in log odds) 57

Odds Ratios using “logistic”

```
.logistic cva age
Logistic regression   Number of obs   =       735
                     LR chi2(1)         =         2.52
                     Prob > chi2        =       0.1127
                     Log likelihood     = -240.98969
                     Pseudo R2         =       0.0051
```

cva	Odds Ratio	StdErr	z	P> z	[95% Conf Int]
age	1.034	.0219	1.59	0.113	.992 1.078

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Standard Error of Odds Ratio

- Logistic regression uses the log odds scale
 - Exponentiate estimates and CI to get inference on odds ratio
- Stata “logistic” provides estimates on odds ratio scale
 - Standard error is from “delta method”
 - CI is from exponentiating log odds CI

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Delta Method Based SE

- In regression models encountered in this class, we can find SE of exponentiated slope parameters

$$\hat{\beta}_1 \sim N(\beta_1, se^2(\hat{\beta}_1))$$

↓ (delta method)

$$e^{\hat{\beta}_1} \sim N\left(e^{\beta_1}, [e^{\beta_1} se(\hat{\beta}_1)]^2\right)$$

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Example: Interpretation

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“From logistic regression analysis, we estimate that for each year difference in age, the odds of stroke is 3.4% higher in the older group, though this estimate is not statistically significant ($P = .113$). A 95% CI suggests that this observation is not unusual if a group that is one year older might have odds of stroke that was anywhere from 0.8% lower or 7.8% higher than the younger group.”

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Logistic Regression and χ^2 Test

-
- Logistic regression with a binary predictor (two groups) corresponds to familiar chi squared test
 - Three possible statistics from logistic regression
 - Wald: The test based on the estimate and SE
 - Score: Corresponds to chi squared test, but not given in Stata output
 - Likelihood ratio test: Can be obtained using post-regression commands in Stata (next quarter)

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Simple Proportional Hazards Regression

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Inference About Hazards

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Right Censored Data

-
- A special type of missing data: the exact value is not always known
 - Some measurements are known exactly
 - Some measurements are only known to exceed some specified value (perhaps different for each subject)
 - Typically represented by two variables
 - An observation time: Time to event or censoring, whichever came first
 - An indicator of event: Tells us which were observed events

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Statistical Methods

- In the presence of censored data, the “usual” descriptive statistics are not appropriate
 - Sample mean, sample median, simple proportions, sample standard deviation should not be used
 - Proper descriptives should be based on Kaplan-Meier estimates
- Similarly, special inferential procedures are needed with censored data

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Notation

Unobserved :

$$\begin{aligned} \text{True times to event : } & \{T_1^0, T_2^0, \dots, T_n^0\} \\ \text{Censoring Times : } & \{C_1, C_2, \dots, C_n\} \end{aligned}$$

Observed data :

$$\begin{aligned} \text{Observation Times : } & T_i = \min(T_i^0, C_i) \\ \text{Event indicators : } & D_i = \begin{cases} 1 & \text{if } T_i = T_i^0 \\ 0 & \text{otherwise} \end{cases} \end{aligned} \quad 66$$

Survival Regression

- There are two fundamental models used to describe the way that some factor might affect time to event
 - Accelerated failure time
 - Proportional Hazards

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Accelerated Failure Time Model

- Assume that a factor causes some subjects to spend their lifetime too fast
 - The basic idea: For every year in a reference group’s lives, the other group “ages” k years
 - E.g.: 1 human year = 7 dog years
 - Ratios of quantiles of survival distributions are constant across two group
 - E.g., report median ratios
 - AFT models include the parametric exponential, Weibull, and lognormal models 68

Proportional Hazards Model

- Considers the instantaneous rate of failure at each time among those subjects who have not failed
 - Proportional hazards assumes that the ratio of these instantaneous failure rates is constant in time between two groups
 - Proportional hazards (Cox) regression treats the survival distribution within a group semiparametrically
 - A semi-parametric model: The hazard ratio is the ⁶⁹ parameter, there is no intercept

AFT vs PH

- Survival analysis: Who does Death prefer?
 - Given a collection of people in a sample:
 - Accelerated failure time models consider how often Death takes somebody
 - If people that Death prefers are available, he/she will come more often
 - Proportional hazards models just compare which people Death chooses relative to their frequency in the population
 - Why is it that Death tends to choose the very old despite the fact that they are less than 1% of the population available 70

Proportional Hazards Model

- Ignores the time that events occur
- Looks at odds of choosing subjects relative to prevalence in the population
 - Can be derived as estimating the odds ratio of an event at each time that an event occurs
 - Proportional hazards model averages the odds ratio across all observed event times
 - If the odds ratio is constant over time between two groups, such an average results in a precise estimate of the hazard ratio 71

Borrowing Information

- Use other groups to make estimates in groups with sparse data
 - Borrows information across predictor groups
 - E.g., 67 and 69 year olds would provide some relevant information about 68 year olds
 - Borrows information over time
 - Relative risk of an event at each time is presumed to be the same under Proportional Hazards 72

Simple PH Regression Model

- “Baseline” hazard function is unspecified
 - Similar to an intercept

Model $\log(\lambda(t | X_i)) = \log(\lambda_{i_0}(t)) + \beta_1 \times X_i$

$X_i = 0$ log hazard at $t = \log(\lambda_0(t))$

$X_i = x$ log hazard at $t = \log(\lambda_0(t)) + \beta_1 \times x$

$X_i = x + 1$ log hazard at $t = \log(\lambda_0(t)) + \beta_1 \times x + \beta_1$

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Model on Hazard scale

- Exponentiating parameters

Model $\lambda(t | X_i) = \lambda_0(t) \times e^{\beta_1 \times X_i}$

$X_i = 0$ hazard at $t = \lambda_0(t)$

$X_i = x$ hazard at $t = \lambda_0(t) \times e^{\beta_1 \times x}$

$X_i = x + 1$ hazard at $t = \lambda_0(t) \times e^{\beta_1 \times x} \times e^{\beta_1}$

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Interpretation of the Model

- No intercept
 - Generally do not look at baseline hazard
 - But can be estimated
- Slope parameter
 - Hazard ratio between groups differing in the value of the predictor by 1 unit
 - Found by exponentiation of the slope from the proportional hazards regression: $\exp(\beta_1)$

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Relationship to Survival

- Hazard function determines survival function

Hazard $\lambda(t | X_i) = \lambda_0(t) \times e^{\beta_1 \times X_i}$

Cumulative Hzd $\Lambda(t | X_i) = \int_0^t \lambda_0(u) \times e^{\beta_1 \times X_i} du$

Survival Function $S(t | X_i) = e^{-\Lambda(t | X_i)} = [S_0(t)]^{e^{\beta_1 \times X_i}}$

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Stata

- `"stcox obsvar eventvar, [robust]"`
 - Provides regression parameter estimates and inference on the hazard ratio scale
 - Only slope with SE, CI, P values

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Example

- Prognostic value of nadir PSA relative to time in remission
 - PSA data set: 50 men who received hormonal treatment for advanced prostate cancer
 - Followed at least 24 months for clinical progression, but exact time of follow-up varies
 - Nadir PSA: lowest level of serum prostate specific antigen achieved post treatment

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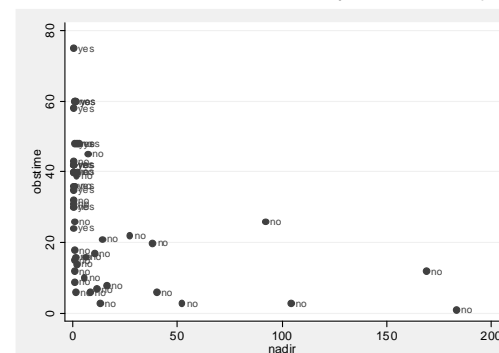
Scatterplots

- Scatterplots of censored data are not scientifically meaningful
 - It is thus better not to generate them unless you do something to indicate the censored data
 - We can label censored data, but we have to remember the true value may be anywhere larger than that

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Obstime vs Nadir (by inrem)

- `scatter obstime nadir, mlabel(inrem)`



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Characterization of Scatterplot

- Outliers
 - ??
- First order trends
 - Certainly downward slope: No censoring at high nadirs
- Second order trends
 - Must be curvilinear (but how much)
- Variability within groups
 - Highest with greater length of observation

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Estimation of Regression Model

```

. stset obstime relapse
. stcox nadir
Cox regression -- Breslow method for ties
No. of subj =      50      No. of obs =      50
No. fail    =      36
Time at risk = 1423

                LR chi2(1) =      11.35
Log lklhood = -113.3      Prob > chi2 =      0.0008
-----+-----
   t | HzRat StdErr   z   P>|z|   [95% Conf Int]
-----+-----
nadir | 1.016  .0038  4.10  0.000   1.008   1.023
    
```

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Interpretation of Stata Output

- Scientific interpretation of the slope

Hazard ratio = $1.015^{\Delta nadir}$

 - Estimated hazard ratio for two groups differing by 1 in nadir PSA is found by exponentiation slope (Stata only reports the hazard ratio):
 - Group one unit higher has instantaneous event rate 1.015 times higher (1.5% higher)
 - Group 10 units higher has instantaneous event rate $1.015^{10} = 1.162$ times higher (16.2% higher)

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Statistical Validity of Inference

- Inference (CI, P vals) about associations requires three general assumptions
 - Assumptions about approximate normal distribution for parameter estimates
 - Assumptions about independence of observations
 - Assumptions about variance of observations within groups

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Normally Distributed Estimates

- Assumptions about approximate normal distribution for parameter estimates
 - Classically or Robust SE:
 - Large sample sizes
 - Definition of “large” depends on underlying probability distribution

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Independence / Dependence

- Assumptions about independence of observations for linear regression
 - Classically:
 - All observations are independent
 - Robust standard error estimates:
 - Allow correlated observations within identified clusters

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Within Group Variance

- Assumptions about variance of response within groups for proportional hazards regression
 - Classically:
 - Mean variance relationship for binary data
 - Proportional hazards considers odds of event at every time
 - Need proportional hazards and linearity of predictor
 - Robust standard error estimates:
 - Allow unequal variances across groups
 - (Do not need proportional hazards or linearity)

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Linearity of Model

- Assumption about adequacy of linear model for prediction of group odds of response with logistic regression
 - The log hazard ratio across groups is linear in the modeled predictor
 - (We can model transformations of the measured predictor)

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Prediction

- We rarely make inference about within group survival probabilities using the proportional hazards model
 - We sometimes use estimated survival curves descriptively
 - Use estimates of baseline survival function
 - Exponentiate the baseline survival to find survival curve for specific covariates

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Relationship to Survival

- Hazard function determines survival function

Hazard $\lambda(t | X_i) = \lambda_0(t) \times e^{\beta_1 \times X_i}$

Cumulative Hzd $\Lambda(t | X_i) = \int_0^t \lambda_0(u) \times e^{\beta_1 \times X_i} du$

Survival Function $S(t | X_i) = e^{-\Lambda(t | X_i)} = [S_0(t)]^{e^{\beta_1 \times X_i}}$

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Implications for Inference

- A hierarchy of null hypotheses
 - Strong (and intermediate) null: Total independence of time to event and X
 - The proportional hazards model holds because the same distribution in every X group
 - Weak null: No linear trend in hazard ratio across X groups

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Classical PH Regression

- Inference about slope tests strong null
 - Tests make inference assuming the null
 - The data can appear nonproportional hazards or nonlinear in log hazard ratio
 - Merely evidence strong null is not true
 - Limitations
 - We cannot be confident that there is a trend in the hazard ratio across groups
 - Valid inference about trend demands correct model

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Robust Standard Errors

- Inference about slope tests weak null
 - Data can appear nonproportional hazards or nonlinear in hazard ratio across groups
 - Robust SE estimates true variability
 - Does not use model based estimates of SE
 - Nonlinearity decreases precision, but inference still valid about first order (linear) trends

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Choice of Inference

- Which inference is correct?
 - Classical PH regression and robust standard error estimates differ in the strength of necessary assumptions
 - As a rule, if all the assumptions of classical PH regression hold, it will be more precise
 - (Hence, we will have greatest precision to detect associations if the linear model is correct)
 - The robust standard error estimates are, however, valid for detection of associations even in those instances

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Interpreting “Positive” Results

- If slope is statistically significant different from 0 using robust SE
 - Observed data is atypical of a setting with no linear trend in hazard ratio across groups
 - Data suggests evidence of a trend toward larger (smaller) hazards in groups having larger values of the predictor

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Interpreting “Negative” Studies

- “Differential diagnosis” of reasons for not rejecting null hypothesis of zero slope
 - There may be no association
 - There may be an association but not in the parameter considered (i.e, the odds of response)
 - There may be an association, but the best fitting line has a zero slope (a curvilinear association in the parameter)
 - There may be a first order trend in the log hazard ratio, but we lacked statistical precision to be confident that it truly exists (type II error)

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Estimation of Regression Model

```

.....
. stset obstime relapse, robust
. stcox nadir
Cox regression -- Breslow method for ties
No. of subj =      50      No. of obs =      50
No. fail    =      36
Time at risk = 1423

LR chi2(1) =      16.79
Log likelihood = -113.3      Prob > chi2 =      0.0000
-----+-----
_+-----+-----+-----+-----+-----+-----+
| t | HzRat StdErr   z   P>|z|   [95% Conf Int]
-----+-----+-----+-----+-----+-----+
nadir | 1.016   .0038   4.10   0.000   1.008   1.023
    
```

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Example: Interpretation

“From proportional hazards regression analysis, we estimate that for each 1 ng/ml unit difference in nadir PSA, the risk of relapse is 1.6% higher in the group with the higher nadir. This estimate is highly statistically significant ($P < .001$). A 95% CI suggests that this observation is not unusual if a group that has a 1 ng/ml higher nadir might have risk of relapse that was anywhere from 0.8% higher to 2.3% higher than the group with the lower nadir.”

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Log Transformed NadirPSA

- Based on prior experience
 - A constant difference in PSA would not be expected to confer same increase in risk
 - Comparing 4 ng/ml to 10 ng/ml is not the same as comparing 104 ng/ml to 110 ng/ml
 - A multiplicative effect on risk might be better
 - Same increase in risk for each doubling of nadir
 - Use log transformed nadir PSA

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Estimation of Regression Model

```

.....
. generate lnadir = log(nadir)
. stcox lnadir, robust
Cox regression -- Breslow method for ties
No. of subj =      50      No. of obs =      50
No. fail    =      36
Time at risk = 1423

LR chi2(1) =      34.04
Log likelihood = -107.3      Prob > chi2 =      0.0000
-----+-----
_+-----+-----+-----+-----+-----+
| t | HzRat StdErr   z   P>|z|   [95% Conf Int]
-----+-----+-----+-----+-----+-----+
lnadir | 1.54   .113   5.83   0.000   1.33   1.77
    
```

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Interpretation of Parameters

.....

- Hazard ratio is 1.54 for an e-fold difference in nadir PSA
 - $e = 2.7183$
- I can more easily understand doubling, tripling, 5-fold, 10-fold increases
 - For doubling: HR : $1.54^{\log(2)} = 1.35$

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PH Regression and Logrank Test

.....

- Proportional hazards regression with a binary predictor (two groups) corresponds to the logrank test
 - Three possible statistics from proportional hazards regression
 - Wald: The test based on the estimate and SE
 - Score: Corresponds to logrank test, but not given in Stata output
 - Likelihood ratio test: Can be obtained using post-regression commands in Stata (next quarter)

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